

# Gender and Leadership in Organisations: The Threat of Backlash\*

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## Abstract

Decisions made by leaders please some people and upset others. We examine whether the possibility of backlash has a differential impact on men's and women's self-selection into leadership roles, and their decisions as leaders. In a laboratory experiment that simulates corporate decision-making, we find that women are significantly less likely to self-select into a leadership position when they can receive backlash. Once in a leadership role, women get more backlash. There are some gender differences in leaders' decision-making and communication styles under the threat of backlash, but little difference in final outcomes. An online experiment sheds light on possible mechanisms.

JEL Codes: C92, D91, J16

Key words: Gender differences, leadership, backlash, experiment.

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# 1 Introduction

Individuals in a position of leadership have to make decisions knowing that some people will be negatively affected by their choices and will therefore be unhappy. In any organisation, for instance, managerial decisions concerning promotions, salary raises, demotions and dismissals necessarily generate negative judgement and possibly anger among some employees. Numerous studies have shown that, holding performance constant, women in a position of power are judged more negatively than men. This is true in politics, business, academia, as well as laboratory and online experiments (Abel, 2022; Ayalew et al., 2021; Boring, 2017; Daniele et al., 2023; Egan et al., 2022; Grossman et al., 2019; Shurchkov and van Geen, 2019).<sup>1</sup> There is also evidence that women are more sensitive to public scrutiny (Alan et al., 2020; Jones and Linardi, 2014) and negative feedback (Mayo et al., 2012; Johnson and Helgeson, 2002), and are more severely discouraged by negative feedback in male-dominated fields or stereotypically male tasks (Ellison and Swanson, 2023; Kugler et al., 2021).

In this paper, we ask whether and to what extent the possibility of backlash from subordinates induces women, more than men, to select out of leadership positions, and to perform differently when in a leadership role. We employ a laboratory experiment that simulates a business environment where managerial decisions, such as rank, job title or task allocations, affect the earnings of some workers positively and the earnings of others negatively.<sup>2</sup> This implies that backlash from unhappy employees, in the form of, at the minimum, negative judgement and disapproving messages, is likely.

In the experiment, a manager is matched with the same two employees for multiple decision-making rounds. This is to simulate the repeated interactions that occur within a team whose members differ in their job and leadership statuses. Following an initial real effort task, the manager's main responsibility is to decide which employee will be high rank and which employee will be low rank, where ranks determine employees' earnings. There can only be one high-rank and one low-rank employee; therefore, rank allocation necessarily creates income inequality between workers. This way, the experiment simulates an environment where the leader/manager has control over some financially meaningful aspects of his or her subordinates' jobs, e.g., their wages, job titles, or yearly bonuses. The manager needs to make the

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<sup>1</sup>See also: Branton et al. (2018); Rheault et al. (2019); Elsesser and Lever (2011); Hengel (2022); Mengel et al. (2017); Sarsons (2017).

<sup>2</sup>This is a departure from the existing experimental studies of leadership, which have typically employed sequential public goods games or coordination games where leaders can induce followers to increase their contributions through leading by example ( e.g., Güth et al., 2007; Grossman et al., 2015; Jack and Recalde, 2015) or through the use of messages suggesting contributions ( e.g., Brandts and Cooper, 2007; Reuben and Timko, 2018). Other important studies of leadership have employed minimum-effort games or real effort tasks where leaders incentivise (Shurchkov and van Geen, 2019) or suggest the effort to be put in by followers (Erkal et al., 2022), or tasks that require leaders to make decisions on behalf of their group (Alan et al., 2020; Born et al., 2022; Reuben et al., 2012).

same decision - for instance whom to assign a bonus to - involving the same team members periodically, e.g. at the end of each fiscal year. Hence, in the experiment, at the beginning of each time period, the manager has to decide whether he or she wants to keep the workers' ranks or switch ranks, i.e., promote the employee who was previously assigned the low rank (and low salary). Importantly, workers engage in a real effort task in every time period, and the manager receives money only from the product of the worker who is assigned the high rank. This generates incentives to assign ranks based on the workers' relative productivity.

In our baseline condition (*No Choice*), we assign the managerial role based on performance in the preceding real effort task. In our treatments, we allow subjects to self-select into the managerial position (*Choice* treatment), and we progressively add the necessity to communicate with employees when assigning ranks (*Choice & Talk* treatment) and the possibility for employees to talk back and send angry emojis to express their disapproval of the rank allocation (*Choice & Backlash* treatment).<sup>3</sup> The use of angry emojis, in particular, allows us to quantify backlash towards managers in the experiment. While, in the field, backlash could take more severe forms, such as sabotage, formal complaints, verbal and physical abuse, the use of emojis to signal disapproval holds significance for two primary reasons. First, with the widespread adoption of social networking sites, individuals, and especially the younger generations, rely heavily on non-verbal cues like emojis to convey their emotions and viewpoints. Therefore, the negative implications associated with the angry emojis in our study are readily comprehensible and salient to participants. Second, if we find that the threat of this form of backlash affects men and women differently, it would suggest that larger gender differences exist for more severe forms of backlash.

Through our treatment manipulations, we are able to isolate the impact of the possibility of backlash on men's and women's self-selection into leadership. Through the analysis of rank allocations and of the messages that managers send to their employees, we are also able to examine whether men and women make different decisions when in a leadership role, conditional on the possibility of backlash. Finally, by examining the angry emojis sent to male and female managers in the *Choice & Backlash* treatment, we can test whether women receive more backlash than men.

We find evidence of a gender gap in self-selection into leadership only when leaders face the possibility of backlash. Neither being responsible for generating income inequality between employees nor having to talk to employees when assigning ranks discourage women, as long as employees cannot send angry messages back. When backlash is possible, men are unaffected, while women are significantly less likely to want to be managers. In line with the existing literature, we find that female leaders get more backlash than male leaders.

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<sup>3</sup>The manager is selected among those who want to be managers based on performance in the preceding real effort task.

When looking at the decisions made by men and women once in a leadership position, we find that the possibility of backlash increases the likelihood of rank-switching for women, but not men. However, this does not lead to profit losses, i.e., we do not see any significant differences in final outcomes across treatments and by gender. Finally, the analysis of the messages sent by managers to workers in the *Choice & Backlash* treatment suggests, in line with the existing literature, that men and women have different communication styles. Female managers praise employees more and use cordial words like thank you and sorry, whereas male managers are more likely to explain the ranking decision and to foster competition among workers. However, these differences only emerge when workers can talk back to managers, which suggests that they reflect men’s and women’s different strategies aimed at minimizing worker backlash.

In order to investigate possible mechanisms behind the observed differences in men’s and women’s responsiveness to the possibility of worker backlash, we conducted an additional experiment on the online platform Prolific. The online experiment allows us to: 1) test whether men and women significantly differ in their aversion to negative feedback; and 2) test whether men and women differ in their beliefs regarding the backlash received by men and women managers in the laboratory experiment. This way, we aim to assess whether women’s higher responsiveness to backlash at the self-selection stage and the rank-switching stage in the laboratory experiment are due to women’s expectations of more severe backlash, as compared to men, and/or more aversion to any backlash. The data show that women score significantly higher in the Brief Fear of Negative Evaluation (BFNE) psychological index (Carleton et al., 2011; Rodebaugh et al., 2004), which indicates that they are significantly more averse to the possibility of negative judgement than men. We also found that, while women (but not men) expect female managers to get more backlash than male managers, there are no statistical significant differences in the likelihood and severity of backlash that men and women expect their own gender to receive. This suggests that the higher responsiveness of women to the possibility of backlash in the laboratory experiment is likely driven by higher sensitivity to (any) backlash.

Our study contributes to the growing literature on the gender leadership gap, which has been documented in all spheres of life. For instance, only 19 percent of firms worldwide have female top managers and only 6 percent of CEOs at S&P 500 companies are women.<sup>4</sup> In politics, women hold only 23 percent of seats in national parliaments worldwide.<sup>5</sup> In academia, averaging across all fields, less than one third of full professors are women. The existing literature has identified behavioural or preference-based constraints to women’s self-

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<sup>4</sup>For recent statistics on the gender leadership gap in the US, see Warner and Corley (2017).

<sup>5</sup>World Bank DataBank: <https://data.worldbank.org/indicator/SG.GEN.PARL.ZS>

selection into top leadership roles.<sup>6</sup> These include risk aversion (see, e.g., Eckel and Grossman, 2008), reticence to initiate negotiations (e.g., Bowles et al., 2007; Babcock and Laschever, 2009; Exley et al., 2020), aversion to competitive environments (e.g., Gneezy et al., 2003; Niederle and Vesterlund, 2007; Flory et al., 2014; Preece and Stoddard, 2015),<sup>7</sup> preferences over job attributes (Wiswall and Zafar, 2017), time spent on low promotability tasks (Babcock et al., 2017), self-stereotyping (Coffman, 2014) and lower willingness to self-promote (Exley and Kessler, 2022).<sup>8</sup> We add to this literature by showing that the threat of backlash contributes to the leadership gap by inducing women, but not men, to select out of leadership roles.

Our study also relates to the small literature in economics on gender differences in leaders' communication styles (Timko, 2017; Manian and Sheth, 2021), which shows that female leaders send less assertive messages than male leaders (Timko, 2017), even when they know that such messages are more effective in impacting others' behaviour (Manian and Sheth, 2021). This is in line with studies in psychology showing that men tend to use more assertive language - e.g., through imperative statements - while women use more affiliative language - e.g., statements of support, agreement, and acknowledgement.<sup>9</sup> We contribute to this literature by showing that gender differences in the language used by male and female leaders to motivate employees do not reflect innate gender differences. Rather, they emerge as different strategies used by male and female managers to minimize backlash.<sup>10</sup> Finally, we add to the growing set of studies providing evidence of significant differences in individual attitudes toward male and female leaders (see, e.g., Abel, 2022; Grossman et al., 2019) by showing that women managers get more backlash than man managers from unhappy employees.<sup>11</sup>

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<sup>6</sup>External demand-side constraints, such as taste-based or statistical discrimination stemming from traditional gender stereotypes concerning men and women's productivities, skills and family constraints, are of course also important. The existing evidence suggests that women are likely to be discriminated against in higher-status jobs, particularly in male-dominated fields. For a review of the literature, see Riach and Rich (2006), Azmat and Petrongolo (2014) and Bertrand and Duflo (2017).

<sup>7</sup>Erkal et al. (2022) show that women are more likely to compete for leadership roles if there is a system in place that, by default, enrol individuals in the competition, while allowing them to opt-out.

<sup>8</sup>Born et al. (2022) also show that being in a male-dominated team reduces women's willingness to lead, due to lower confidence and expectations of team support.

<sup>9</sup>See, for instance, Leaper and Smith (2004) and Kern et al. (2016).

<sup>10</sup>We also indirectly contribute to the literature on gender differences in executives' performance (see, e.g., Adams and Ferreira, 2009; Matsa and Miller, 2014). In line with existing studies, we find that female leaders are more likely to make decisions that benefit low-earning employees, but this holds only if managers have to personally communicate their decisions to their subordinates, and if they can receive angry messages back as a result.

<sup>11</sup>There is also a large literature on attitudes toward male and female leaders in psychology, sociology and management. These studies typically either provide written description of leadership situations, varying the sex of the leader, or use trained actors to lead, allowing the experimenters to control the degree of success the leader achieves (see, e.g., Swim et al., 1989).

## 2 The Experiment

### 2.1 Design

The experiment consists of 6 active stages (Stages 1 to 6), followed by a survey, as shown in Figure 1. An important feature of our design is the method used to reveal subjects' genders to other participants without making gender artificially salient in the game. We achieved this by asking subject to fill in a brief survey at the very beginning of the session, before Stage 1. The survey registers subjects' age, gender, field of study, and previous participation in an experiment. The answer to the gender question leads to a pre-determined list of either male or female names, which we took from Bertrand and Mullainathan (2004)'s correspondence study of race-based discrimination.<sup>12</sup> The male subjects see a list of male names and the female subjects see a list of female names. We inform subjects that for the duration of the experiment they will be identified with a fictitious name, and we invite them to pick a name from the gender-specific list they see on their screen.<sup>13</sup> We do not allow two or more subjects to choose the same name, so each name disappears from the list in real time when picked by another participant.

Stage 1 follows. In this stage, and in the following five stages, subjects engage in a real effort task. Previous studies of leadership have typically used public goods games or coordination games. In order to resemble firm environments where managers are chosen based on qualifications, in our study we wanted a game/task that would allow us to clearly assess participants' relative performance and select the best performing subject in a group as the leader. Specifically, we wanted a gender-neutral task requiring cognitive thinking and focus, where men and women would be equally confident and would perform equally well. We therefore chose a language task, as it has been shown (e.g., Dreber et al., 2014; Niederle, 2016; Shurchkov, 2012) that language-based tasks are less likely than math-based tasks to generate gender differences in both self-confidence and performance in competitive environments.

In Stage 1, our real-effort task consists in finding a 4-letter word in a 6x6 letter matrix in 5 minutes, for a maximum of 20 matrices.<sup>14</sup> Subjects play individually. They receive an

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<sup>12</sup>As our focus is on gender differences, we use Bertrand and Mullainathan (2004)'s list of distinctively white sounding names only. Distinctive names are those that have the highest ratio of frequency in the corresponding racial group.

<sup>13</sup>We did not ask subjects to use their real names as we did not want to lift anonymity nor did we want the potential confounding bias of race, nationality or ethnicity associated with the actual name of the subject to play a role in the experiment.

<sup>14</sup>The decision screen is divided in two halves. On the left, subjects see the matrix and on the right, they see a list of 40 words. Each puzzle contains two words that appear on the list. The screen is reproduced in the Online Appendix. In order to earn money, subjects have to identify one word per puzzle. We used the website <http://tools.atozteacherstuff.com/word-search-maker/wordsearch.php> to create the puzzles and the website <http://www.thefreedictionary.com> to find words of varying lengths. We ran some pilots of the puzzle task with varying levels of difficulty with different sizes of the matrix, different word lengths, and different ways in which words could be identified in the puzzle (forward, backward, up, down, diagonal etc).

endowment of 40 Experimental Currency Units (ECU)<sup>15</sup> and earn 2 ECU for each puzzle they solve correctly in 5 minutes. At the end of Stage 1, subjects receive feedback on their absolute performance and are provided instructions on the following 5 stages (Stages 2 to 6) of the experiment. Crucially, they are randomised into groups of 3 and they are shown the fictitious names of their group members. In order to simulate male-dominated environments, the randomisation algorithm created groups of 2 men and 1 woman, whenever possible.<sup>16</sup>

Since our interest is in work environments where manager and workers interact multiple times, so that a manager is responsible for the long-run income of his or her employees, subjects remain in the same group for the duration of the experiment. In Stages 2 to 6, two group members play in the role of workers and one in the role of manager. The roles of manager and worker are assigned at the end of Stage 1 according to treatment-specific rules, and retained through Stage 6. The manager gets a fixed wage of 100 ECU and his/her main task is to decide, at the beginning of each stage of the experiment, which worker will be rank A and which worker will be rank B in the following stage. The rank A worker gets a wage of 80 ECU, while the rank B worker gets a wage of 20 ECU. After the rank allocation, all members of the group engage in a similar puzzle task as in Stage 1 of the experiment.<sup>17</sup> Each correctly solved puzzle generates 2 ECU in addition to the initial wage. Moreover, each puzzle solved correctly by the Rank A worker generates 2 ECU also to the manager. This is to incentivize the manager to take the workers' relative productivities into account when making the rank allocation decision. In sum, the earnings from each of the 5 active stages (Stages 2 to 6) of the experiment are determined as follows:

- The manager gets 100 ECU + 2 ECU per puzzle + 2 ECU per puzzle solved by the rank A worker;
- The rank A worker gets 80 ECU plus 2 ECU per puzzle;
- The rank B worker gets 20 ECU plus 2 ECU per puzzle.

Stages 3 to 6 are identical to Stage 2. However, at the end of each stage of the experiment, the manager is informed about the performances of the current rank A and rank B workers and has to decide whether to keep or reassign ranks before the next stage begins.

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We found the configuration of finding 4-letter words that appear horizontally or vertically in a 6X6 matrix with a time of 5 minutes in Stage One to be optimal in creating enough heterogeneity in performance among subjects.

<sup>15</sup>Earnings from the experiment were later converted to dollars at the exchange rate of 6 ECU for 1 USD, as explained in Section 2.3.

<sup>16</sup>We ended up having 63% of the groups made of one woman and two men, 30 percent made of two women and one man, and 7 percent made of men only.

<sup>17</sup>In order to account for learning effects, while in Stage 2 we keep the time limit to solve the 20 matrices equal to 5 minutes, we reduce the time to 4 minutes in Stages 3 and 4, and to 3.5 minutes in Stages 5 and 6.

Following Stage 6, subjects fill a post-experiment questionnaire,<sup>18</sup> where we elicit demographics, previous leadership experiences, and answers to personality questions that allow us to generate the Big 5 Agreeableness Personality Index, which has been shown to be significantly higher in women than men (Schmitt et al., 2008) and to correlate negatively with leadership ambition (e.g., Ertac and Gurdal, 2012).<sup>19</sup>

Some features of our design require further discussion. First, the difference between the wages of the manager and the rank A worker may seem small when compared to the wage difference between the rank A and the rank B workers. We note that while the manager’s fixed wage is only 20 percent higher than the fixed wage of the rank A worker, there is also an important difference in the additional earnings generated by the puzzle-solving task. These additional earnings can be thought of as bonus earnings in organisations. The manager bonus is at least twice as large, on average, than the bonus earned by the rank A worker. This is because the manager receives earnings from the puzzles solved by the rank A worker (2 ECU per puzzle) in addition to the puzzles the manager solves (2 ECU per puzzle). Given that the manager tends to be the most productive individual in the group (partly by design), this means that he or she makes at least twice as much as the rank A’s bonus earnings. So, the actual difference between the total earnings of manager and rank A worker is 20 percent in fixed wages and 100 percent in bonus earnings.<sup>20</sup> These differences are not trivial, and are meant to provide meaningful incentives for participants to self-select into the manager position, especially considering that, if they assume the roles of worker, they could end up being the Rank B worker for the payoff-salient stage of the experiment (which is randomly chosen as explained in Section 2.3).

The difference in the earnings of the rank A and rank B workers is large and may seem unrealistic for two employees in the same line of work, who are assigned the same task. However, the large earning gap is necessary in the lab setting both to induce workers to care enough about the ranks and the possibility of a promotion, and to make managers believe

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<sup>18</sup>Due to a computer glitch, we were unable to conduct the post-experiment questionnaire on 12 participants.

<sup>19</sup>The Agreeableness Index measures the tendency to be kind, altruistic, trusting and trustworthy, and cooperative. There is evidence from psychology studies (e.g., Judge and Bono, 2000) that agreeableness predicts transformational leadership, i.e. leadership that operates through inspiration, intellectual stimulation and individual consideration. The index is based on individual answers (on a 5-point agree/disagree Likert scale) to the following questions: 1) I feel little concern for others; 2) I am interested in people; 3) I insult people; 4) I am not interested in other people’s problems; 5) I have a soft heart; 6) I am not really interested in others; 7) I take time out for others; I feel others’ emotions; 8) I make people feel at ease. The index ranges between 0 and 40. For more information on how the index is calculated, see: <https://openpsychometrics.org/printable/big-five-personality-test.pdf>

<sup>20</sup>Assuming average productivity for both manager and rank A worker, i.e., about 14 puzzles, this means a bonus of 28 ECU for the worker and at least 56 ECU for the manager. This means earnings of 108 for the rank A worker and at least 156 for the manager, i.e., the manager makes over 40 percent more than the rank A worker and nearly 200 percent more than the rank B worker (assuming average productivity also for the Rank B worker).



that the workers care about the rank allocation decision. In the field, there are other aspects of the job that affect the significance individuals attach to a promotion, besides the wage. For instance, a promotion affects the career ladder, as it increases the likelihood of future career advancements; it also elevates the individual’s social status among co-workers and the broader community. A promotion may also come with other on-the-job benefits such as more flexible work hours. In the laboratory setting we need to abstract from these features. Hence, the large pay gap is meant to capture some of the benefits associated with a promotion that individuals care about, but that cannot be fully accounted for in our setting.

Finally, in the experiment, the nature of the task is such that the manager is always able to accurately assess the relative performances of the two employees. This allows us to define and clearly measure managerial performance as: 1) the likelihood that the manager makes the first rank allocation based on performance in Stage 1, and 2) the likelihood that the manager makes an ex-post mistake in the rank allocation, by assigning rank A to a worker who ends up not performing as well as the other worker. We acknowledge that, in many settings, workers’ performances cannot be objectively or precisely measured, and a manager’s rank allocation decisions are at least partly discretionary; it is the lack of transparency and the subjectivity of the decision process that may be especially conducive to worker backlash. Note that even though our task generates objective workers’ rankings, we still allow for lack of transparency and perceived subjectivity of the manager’s decisions by not disclosing relative performances to the workers. In other words, the workers do not know how they compare to each other, and do not know what criteria the manager followed to allocate ranks.

## 2.2 Treatments

In our *No Choice* (NC) treatment, at the end of Stage 1, in each group of three participants the manager is chosen based on performance in Stage 1.<sup>21</sup> Recall that subjects participate in the Stage 1 real effort task individually without knowing anything about Stages 2 to 6, and therefore ignoring the fact that their performance would determine their role in the subsequent stages of the experiment. This prevents competition-driven anxiety from playing a role in determining subjects’ performance and subsequent chances of becoming the manager of the group. While there is a large literature documenting gender differences in competitiveness and self-confidence (especially in math-based tasks), we abstract from both factors by design, in order to be able to isolate the role played by the possibility of worker backlash in the origination of gender differences in leadership.

At the beginning of Stage 2, subjects are informed that the manager was chosen based on

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<sup>21</sup>Since, by design, we have more men than women participating in each session of the experiment – due to the objective of having male-dominated groups – in all treatments we break ties in favour of women. Subjects are unaware of this.

performance rather than randomly. This is important, as we aimed to simulate an environment where employees could not doubt the qualifications of their manager. This way, any differences in workers' attitudes toward male versus female managers could not be attributed to differential subjective beliefs about the right of the manager to hold his or her role in the group.

In our *Choice* (C) treatment, we allow subjects to self-select into the leadership position. At the end of Stage 1, after receiving information about the next five stages of the experiment and the fictitious names of their group members, we ask subjects to state whether they would like to be the manager of their group. From the subset of those who want to be manager, we choose the manager based on performance in Stage 1, as in the *No Choice* treatment. Since, by design, the manager is always the highest earner of the group, everybody should want to be manager.<sup>22</sup>

In the *Choice & Talk* (CT) treatment, we inform participants that the manager will have to assign ranks to workers in each stage, as before, and will also have to communicate with each worker when assigning ranks by sending a free-form message to the rank A and the rank B workers before the beginning of the next stage of the real effort task. Subjects are told that the manager can write anything he or she wishes to communicate to each worker. The process is repeated at each rank-allocation stage. As in the *Choice* treatment, after providing information about the next stages of the game, including tasks and payoffs of manager and workers, as well as the fictitious names of their group members, we ask subjects to state whether they want to be the manager of their group. If more than one group member wants to be manager, we select the manager following the same procedure as in the *Choice* treatment. In this treatment, an aversion to having to repeatedly provide feedback to subjects who are affected either positively or negatively by the rank allocation decision, may lead subjects to select out of the manager role<sup>23</sup> and impact men and women differentially.

Finally, in our *Choice & Backlash* (CB) treatment, the manager still has to send a free-form message to the rank A and the rank B workers, yet in this case, each worker can send a free-form message back to the manager. Moreover, the worker assigned rank B can send up to 5 angry emojis to the manager to express disapproval of the rank allocation. The messages sent by the two workers, including the angry emojis, are displayed to the manager before the

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<sup>22</sup>Individuals may perceive the manager selection as a competitive process, hence, it is possible that gender differences in competitiveness (Niederle and Vesterlund, 2007) may hold women back. However, contrary to standard competition tasks, in our setting not becoming manager, i.e., losing the competition, does not come with a loss, since the consequence would be that the subject plays the game as a worker, i.e., same outcome as if he or she did not self-select into leadership. In other words, there is no risk in wanting to be manager. Still, there may be a gender difference in willingness to receive negative feedback about the outcome of the competition, as it could reveal information about one's relative ability in the real effort task. Even if that were true, it would apply to all our self-selection treatments; therefore, it would not act as a confound in our study of the behavioural mechanism of interest.

<sup>23</sup>For instance, subjects may feel uncomfortable at the thought of having to talk to workers assigned to be rank B, possibly multiple times during the course of the experiment.

next real-effort task begins. The process is repeated at each rank-allocation stage. As before, participants receive information about the rules governing Stages 2 to 6 of the experiment, including the presence of two-way communication and the possibility of receiving angry emojis from rank B workers, before they are asked whether they would like to be the manager of their group.

The *Choice & Backlash* treatment most closely simulates a work environment where managers have to communicate their promotion or task allocation decisions to employees, knowing that employees will be able to talk back and express their discontent. By comparing men’s and women’s behaviours in this treatment versus the *Choice* and the *Choice & Talk* treatments we are able to identify the role of anticipation of worker backlash on self-selection into a leadership role. Note that the decision to use angry emojis in the *Choice & Backlash* treatment follows the aim to make both workers and manager aware that expression of disapproval of the manager’s decision is possible and allowed in the experiment - which is not easy to achieve with the use of free-text messages only, and without framing the instructions in a way that may prime or lead subjects to say specific things to each other. Moreover, emojis have become a widespread form of communication among the younger generations, and the meaning of an angry emoji is universally understood. Therefore, the use of emojis allows us to: 1) hold the boundaries of potential worker discontent constant (between 0 and 5); 2) quantify the attitudes of workers toward managers; and 3) examine how male and female managers respond to angry messages from employees, holding such messages constant.<sup>24</sup> We acknowledge that angry emojis do not capture more severe forms of backlash that may take place in the workplace, such as formal complaints, sabotage, and verbal or physical abuse. However, we argue that they can still provide informative insights into participants’ sentiments and reactions, especially given their widespread use in online communication to convey dissatisfaction, frustration, or disapproval. Evidence of a gender difference in the reaction to this form of backlash would indicate that a larger difference exists for more severe forms of backlash.

In addition to self-selection into a manager role, our design allows us to investigate gender differences in managers’ decision-making. By design, only the performance of the rank A worker generates additional earnings to the manager. Therefore, the manager should assign rank A to the worker that he or she expects to be the best performer in the upcoming stage. The only information available to the manager is the workers’ past performances, so we expect the difference in the workers’ past productivities to play a primary role in the rank allocation decisions. However, managers’ distributional concerns and/or the desire to avoid the discomfort of having to tell the same worker that he or she has been assigned rank B multiple times,

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<sup>24</sup>We do not allow rank A employees to send happy emojis. Although this could be an interested extension of our study, as we discuss further in the Conclusion, the focus here is the anticipation of negative judgement and hostile messages from unhappy employees, hence the use of angry emojis only.

and/or the objective to minimise the receipt of angry emojis, may lead managers to switch ranks between the two workers over the course of the experiment in order to equalise their earnings. Our treatment comparison allows us to examine gender differences in managers' rank switching behaviour, and the mechanisms leading to them. We are also able to assess whether differential rank switching propensities lead to different outcomes, i.e., mistakes in rank allocations and profit losses.

Finally, by examining the free-form messages sent by managers to workers in the *Choice & Talk* and in the *Choice & Backlash* treatments, not only can we study gender differences in communication styles, but we can also examine the extent to which the languages used by male and female managers change under the threat of worker backlash. Furthermore, the analysis of both free-form messages and angry emojis sent by workers to managers provides a comprehensive assessment of differences in workers' attitudes toward male and female managers.

### 2.3 Implementation

We conducted 29 experimental sessions at the Laboratory for Research In Experimental Economics (LREE) at Southern Methodist University. We involved a total of 417 participants, of which 41 percent are women, as shown in Table 1. We employed a between-subject design, with each subject participating in only one session and one treatment. In each session, we had multiple independent groups of one manager and two workers. Groups were fixed for the duration of the experiment, and members of each group made decisions independently from all the other groups participating in a session.

As described in Section 2.2, the experiment consisted of an initial brief survey and name-assignment stage, followed by six active stages plus a post-experiment survey. Subjects were presented with the instructions for each stage on their computer screen immediately before that stage began. Additionally, at the beginning of Stage 1, they received hand-outs and verbal instructions about the puzzle-solving task, and at the end of Stage 1 they received further verbal and written instructions (hand-outs) about the rules applying to Stages 2 to 6 of the experiment.<sup>25</sup>

Only one randomly selected active stage of the experiment was used for actual payments. Experimental earnings were converted from ECU to dollars at the exchange rate of \$1 for 6 ECU. The experiment was programmed in z-Tree (Fischbacher, 2007) and subjects were recruited among pre-registered LREE students.

In order to guarantee anonymity, at the beginning of each session subjects were randomly assigned an identification number, which they kept for the duration of the experiment. At no

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<sup>25</sup>The experiment instructions are provided in the Online Appendix.

point during the experiment did we ask subjects to reveal their names and, although actual names were used during the payment process for accounting purposes, we informed subjects that we would not register their names and therefore would not be able to link them to the choices made in the experiment. Each session lasted between 60 and 90 minutes, with average earnings of \$28 per subject, including a \$10 show-up fee.

## 2.4 Estimation Strategy

We test the effects of the treatments on men’s and women’s decision to become a manager by estimating equation 1 below, where we pool the men and women samples and include interactions between a gender dummy and each treatment:

$$Y_i = \beta_0 + \beta_1 CT_i + \beta_2 CB_i + \beta_3 F_i * CT_i + \beta_4 F_i * CB_i + \gamma F_i + \delta \mathbf{X}_i + u_i \quad (1)$$

$Y_i$  is participant  $i$ ’s stated desire to be the manager of the group,  $CT_i$  is a dummy equal to 1 if  $i$  is in the *Choice & Talk* treatment and  $CB_i$  is a dummy equal to 1 if  $i$  is in the *Choice & Backlash* treatment. The *Choice* treatment is the excluded category.  $F_i$  is a dummy variable equal to 1 if subject  $i$  is a woman. In this specification,  $\beta_1$  and  $\beta_2$  indicate the impacts of the *Choice & Talk* and *Choice & Backlash* treatments on men, while  $\beta_3$  and  $\beta_4$  capture the differential impacts of the two treatments on women. We also report p-values generated by Wald tests for linear combinations of the estimated coefficients of interest, e.g. for  $\beta_1 + \beta_3 = 0$ , to assess whether the treatments had a significant impact on women. In our most comprehensive specification, we include  $X$ , a vector of demographic controls, i.e., age, field of study, being a native English speaker, having held a leadership position outside the lab and the Big 5 agreeableness measure. We also control for the performance in Stage 1 of the experiment. Even though we employ a between-subject design and groups are independent from each other, the gender composition of the group may affect self-selection into leadership, as recently shown by Born et al. (2022); therefore, we control for the number of women in the group and we cluster the standard errors at the group level. We estimate equation (1) using linear probability models, since interpreting interaction terms in non-linear models is not straightforward, as the marginal effect of an interaction term may not be the same as the estimated coefficient, and further the standard t-test is inaccurate (Norton et al., 2004).<sup>26</sup>

We adopt the same estimation strategy when examining gender difference in managers’ decision-making. However, this analysis also includes the *No Choice* treatment and explores the fact that managers made active decisions affecting the workers in their group multiple times. In these regressions, we still keep the *Choice* treatment as our benchmark category, as

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<sup>26</sup>We also estimate treatment effects for men and for women separately, by splitting the sample by gender, and report the results in the online appendix.

shown in equation 2:

$$Y_{it} = \beta_0 + \beta_1 NC_i + \beta_2 CT_i + \beta_3 CB_i + \beta_4 F_i * NC_i + \beta_5 F_i * CT_i + \beta_6 F_i * CB_i + \gamma F_i + \eta DP_{it-1} + \delta \mathbf{X}_i + \lambda_t + u_{it} \quad (2)$$

where  $Y_{it}$  is manager  $i$ 's decision in Stage  $t$ .  $NC_i$  is a dummy equal to 1 if the manager is in the *No Choice* treatment,  $CT_i$  is a dummy equal to 1 if the manager is in the *Choice & Talk* treatment,  $CB_i$  is a dummy equal to 1 if the manager is in the *Choice & Backlash* treatment. Here, we control for the difference in the performances of the two workers assigned to manager  $i$  in the previous Stage,  $DP_{it-1}$ , as it is likely to affect manager decision-making, and we include stage fixed effects. As before, we cluster the standard errors at the group level, which in this case is the equivalent of clustering the standard errors at the individual level, as each group has the same manager for the duration of the experiment. Since the leadership experiment starts at  $t = 2$ , and we control for the rank status and performance of the worker in the previous round, the analysis is restricted to  $t > 2$ . We employ three measures of managerial decision-making: 1) whether the manager switches worker ranks after the first rank allocation; 2) Whether the manager makes an ex-post mistake in the rank allocation by promoting a worker who ends up performing worse than the other worker, and 3) the amount of profits that are lost due to mistaken rank allocations.

Given that we have four outcome variables measuring individuals' decision-making as a result of our treatments, hence multiple hypotheses, we correct the p-values associated to individual hypotheses by employing the step-down multiple testing method developed by Romano and Wolf (2005). We present the results for the split samples, including the multiple hypothesis correction, in the online appendix. We report estimates from the full sample with interaction terms (equations 1 and 2) in the main text.

When analysing the number of angry emojis that rank B workers send to male versus female managers in the *Choice & Backlash* treatment, we estimate equation 3 below,

$$A_{jit} = \beta_0 + \beta_1 FM_i + \beta_2 P_{jt-1} + \beta_3 RankB_{jt-1} + \delta \mathbf{X}_j + \lambda_t + u_{jit} \quad (3)$$

where  $A_{jit}$  is the number of angry emojis, including 0, that the rank B worker  $j$  sent to manager  $i$  at the beginning of Stage  $t$ ,  $FM_i$  is a dummy equal to 1 if manager  $i$  is a woman,  $P_{jt-1}$  is the performance of worker  $j$  in the task in Stage  $t - 1$ , and  $RankB_{jt-1}$  is a dummy equal to 1 if worker  $j$  was rank B also in Stage  $t - 1$ . As before, the analysis is restricted to  $t > 2$ . We include our set of controls, stage fixed effects and we cluster the standard errors at the group level. To account for the fact that rank B workers can send angry emojis *after* reading the free-form message sent by their manager, we categorise the manager messages

based on their content and in our most comprehensive specification we also control for the type of message sent by the manager.

Finally, in order to test whether male and female managers are more or less likely to promote a rank B worker after receiving angry emojis, we restrict the analysis to the *Choice* & *Backlash* treatment and estimate equation 4:

$$Y_{it} = \beta_0 + \beta_1 A_{it-1} + \beta_2 F_i * A_{it-1} + \beta_3 DP_{it-1} + \gamma F_i + \delta \mathbf{X}_i + \lambda_t + u_{it} \quad (4)$$

where  $Y_{it}$  is a dummy equal to 1 if manager  $i$ , at the beginning of Stage  $t$  promotes the worker who was rank B in Stage  $t - 1$ ,  $F_i$  is a dummy equal to 1 if manager  $i$  is a woman,  $A_{it-1}$  is the number of angry emojis that  $i$  received from the rank B worker when allocating ranks at  $t - 1$ ,  $DP_{t-1}$  is the difference between the performance of the rank A and the rank B workers at  $t - 1$ . Since managers can first receive backlash at  $t = 2$  the analysis of rank allocations in reaction to backlash focuses on  $t > 2$ . In order to test whether male and female managers respond differently to the receipt of angry emojis, we include the interaction between the gender of the manager and the number of angry emojis received from the rank B worker at  $t - 1$ . As before, we include stage fixed effects and cluster the standard errors at the group level.

### 3 Results

We start by describing our subject pool and conducting balance tests across treatments. A total of 245 men and 172 women participated in the experiment. In Table A1, in the Appendix, we report descriptive statistics for our male and female sample pools, i.e. their average age, whether they were majoring in STEM, Business or Economics or in a different field, whether they were native English speakers, whether they reported having held a leadership position, and their average Big 5 Agreeableness Index. The average age is 22.9, with no significant gender differences across treatments and genders. Most of our participants are STEM, Business or Economics majors, although the percentage of men majoring in these fields (83 percent) is significantly higher than the percentage of women (64 percent,  $p = 0.000$ ). The percentage of men studying STEM or Economics is higher in the *Choice* treatment than in the *Choice* & *Talk* ( $p = 0.044$ ) treatment. There are no significant differences across treatments for female participants. Most subjects reported having held a leadership role in the past, with no significant differences across treatments and genders. About 41 percent of men and 55 percent of women are native English speakers, ( $p = 0.008$ ). In line with the existing literature, our female participants score significantly higher than men in the Big 5 Agreeableness Index ( $p = 0.004$ ), with no significant differences across treatments for both the male and the female



samples. The only individual characteristic that is not balanced across treatments is the likelihood for a subject to be a native English speaker. This likelihood is higher in the *Choice & Talk* treatment than in the other treatments.

Before presenting and discussing our main findings, we assess possible gender differences in the performance in the real effort task employed in the study. Recall that we aimed to design a real effort task that would be as much as possible gender-neutral. On average, subjects solved 14 out of 20 puzzles correctly in Stage 1. The only statistically significant difference we see across genders and treatments is that women are better at the task in the *Choice & Talk* treatment than in the other treatments ( $p = 0.04$ ). Figure A1 in the appendix, shows the distribution of puzzles solved correctly by male and female workers over the five stages of the experiment, averaging across all treatments. The distributions are remarkably similar. Table A2 in the appendix shows estimates from regression analysis where the dependent variable is the number of correctly solved puzzles per stage. We find no statistically significant gender differences within and across treatments, which suggests that we were successful in employing a gender-neutral task.

We also wanted our real effort task to lead to stable performance rankings within a group. In other words, we wanted to employ a task whereby being the best(worst) performer in the task in Stage 1 would be a good predictor of the likelihood of being at the top(bottom) of the group in the subsequent stages. This is what we see in the data, i.e., if a subject is the top(bottom) performer in Stage 1, he or she is the top(bottom) performer in about 80 percent of the following stages. When restricting the analysis to the two workers, the best performer in Stage 1 is the best performer in 77 percent of the following stages, i.e., about 4 of the following 5 stages.

In what follows, we present and discuss the core results of the paper, i.e., the effects of our treatments on men’s and women’s self-selection into the manager position (Section 3.1), and on their decisions once in the manager role (Section 3.2). We then present evidence of the backlash received by men and women managers, and of their reactions to such backlash (Section 3.3). We conclude by conducting robustness checks on our main findings (Section 3.4).

### 3.1 Self-selection into the leadership role

Almost all subjects wanted to be a manager in both our *Choice* and our *Choice & Talk* treatments, with no significant differences between men and women, as shown in Table 2 and Figure 2. However, a large and statistically significant gender gap emerges in our *Choice & Backlash* treatment, where we see 78 percent of women self-select into the manager role as opposed to 95 percent of men ( $p = 0.007$ ). The decline in managerial ambition among women



is significant compared to both the *Choice* ( $p = 0.055$ ) and the *Choice & Talk* treatments ( $p = 0.011$ ). On the other hand, we do not see a decline in self-selection into the manager role in the *Choice & Talk* as compared to the *Choice* treatment. This suggests that, absent the possibility of worker backlash, women are not less willing than men to assume the leadership role in our setting.

This is confirmed by the estimates displayed in Table 3.<sup>27</sup> We start by testing for treatment effects by including our treatment variables only. We then gradually add interaction terms and controls, i.e., individuals' performance in the real effort task in Stage 1, and a dummy equal to 1 if there were two women in the group. Note that for 12 subjects we have missing values for three control variables: past leadership positions, being a native English speaker and the Big 5 Agreeableness index. This is due to a software glitch that prevented us from conducting the post-experiment survey in one session of the *Choice* treatment. We follow standard practice and input the missing values using the variable mean for the Big 5 Agreeableness index, which is a continuous variable. For the other two controls, which are binary variables, we input the most common values. We also include a dummy variable equal to 1 for the 12 subjects that had inputted values for the three missing control variables.<sup>28</sup>

The estimates show evidence of a 16 percentage point gender leadership gap in the *Choice & Backlash* treatment only. Women, but not men, are less likely to self-select into the manager role only when facing the possibility of backlash (p-value=0.047, test for the sum of coefficients of *Choice Backlash* and its interaction with the female dummy in Column 2, and p-value=0.069 in column 4). Importantly, only for women, the likelihood of self-selection is also significantly lower in the *Choice & Backlash* treatment when compared to the *Choice & Talk* treatment (p-value=0.013), suggesting that it is the possibility of backlash and not the necessity to give feedback to employees that deters women from wanting to be managers.

Among the controls, we find that performance in the task is a strong predictor of (only) women's managerial ambition. We therefore examine whether gender differences in the decision to be a manager are more extreme for bottom performers than for the middle and top performers. Figure A2 in the appendix reports the percentages of individuals who want to be managers among men and women in all treatments by performance terciles.<sup>29</sup> In the *Choice* and *Choice & Talk* treatments, we do not see evidence of significant gender differences in self-selection into the manager role for any tercile. This is to be expected given that over 90 percent of men and women self-selected into leadership in these treatments. In the *Choice &*

<sup>27</sup>Estimates obtained for the male and female samples separately are shown in Table A3 in the appendix.

<sup>28</sup>In the working paper version of this manuscript, we had dropped the 12 subjects with missing values from the analysis reported in column 4 of Table 3. The resulting estimated coefficients were very similar, both in terms of magnitude and statistical significance.

<sup>29</sup>The bottom tercile is made of subjects who completed 12 or less puzzles correctly in Stage 1 of the experiment. The middle tercile consists of students who completed more than 12 puzzles but less than 20 puzzled in Stage 1. The top tercile is made of students who completed 20 out of 20 puzzles in Stage 1.

*Backlash* treatment, significant gender differences in volunteering exist among both bottom ( $p = 0.065$ ) and middle performers ( $p = 0.009$ ). Importantly, the gender gap closes among the very top performers ( $p = 0.883$ ), i.e. the men and women who solved 20 out of 20 puzzles in Stage 1. The figure also shows that while there are no significant differences in self-selection among performance terciles in the male sample for either treatment, in the *Choice & Backlash* treatment, female top performers are significantly more likely to volunteer than female bottom performers ( $p = 0.058$ ). This suggests that, contrary to men, women feel that, in order to be managers of their group in a setting when employees can express negative feedback, they need to be extremely good at the task.

The only other robust determinant of self-selection into leadership for women only is the subjects' field of study. Women majoring in STEM, Business and Economics are more likely to want to be managers. This indicates that women who have already self-selected into fields that are typically male-dominated and conducive to competitive high-paying jobs, are also more likely to self-select into leadership roles in the experiment. The gender gap in leadership is larger for women who are not majoring in STEM, business or economics, although the data are noisy since only 25 percent of the sample belongs to this category.<sup>30</sup>

### 3.2 Managers' decision-making

We start by examining the likelihood that the first rank assignment, at the beginning of Stage 2, is based on workers' relative productivity, i.e., the number of correctly solved puzzles in the previous stage. Overall, 94 percent of the managers assigned rank A to the best performing worker in Stage 1,<sup>31</sup> with no statistically significant differences by manager gender and/or by treatment.

Next, we examine managers' tendency to switch worker ranks across stages. As previously noted, we designed the task with the aim of having workers' relative performances be stable over time. On average, the best worker in stage 1 outperforms the other worker in nearly 4 of the 5 subsequent stages. Therefore, once ranks are established based on past performance at the beginning of Stage 2, switching ranks and promoting the rank B worker may generate less money than if the ranks had been kept unchanged. However, a desire to equalise workers' payoffs, or to reward the rank B worker if he or she performs equally or better than the rank A worker in a given stage, or to reduce the likelihood of backlash, may induce managers to switch ranks across stages.

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<sup>30</sup>In the male sample, the only significant variable is the Big 5 Agreeableness score, which, in line with the existing studies, appears with a negative sign, suggesting that less agreeable, hence more competitive and aggressive men, are more likely to self-select into the manager role.

<sup>31</sup>If the two workers solved the same number of puzzles, we code our outcome variable as a 1, no matter the worker who was chosen to be rank A.

Overall, averaging all treatments, male and female managers switch ranks 1.03 and 1.44 times, respectively ( $p\text{-value}=0.053$ ). Workers' relative past performance is a strong predictor of a manager's decision to switch ranks. Rank B workers who outperform rank A workers are promoted 56 percent and 65 percent of the times by male and female managers respectively ( $p\text{-value}=0.215$ ). We estimate equation 2 of Section 2.4 and report the results in columns 1 to 3 of Table 4.<sup>32</sup> Among the controls, in addition to those employed in Table 3, we include the difference in the performances of the two workers, which is a strong predictor of manager ranking decisions. The estimates show that both men and women are more likely to switch ranks when they have to talk to workers at the rank allocation stage, in the *Choice & Talk* treatment. The estimated total treatment effects suggest that having to communicate with workers when assigning ranks increases male and female managers' likelihood of switching ranks by about 18 percentage points (Column 3). Such likelihood increases further for female managers by about 12 percentage points under the possibility of worker backlash.

In columns 4 to 6 of Table 4 we test whether the treatments have a significant impact on men's and women's likelihood of assigning rank A to a worker who ends up performing worse than the other worker - we call this an ex post mistake in the rank allocation. We find that the possibility of backlash leads women to make more ranking mistakes as compared to men. However, this is not due to an increase in women's likelihood of making mistakes as a result of the possibility of backlash. The p-value associated with the test of equality of the impacts of *Choice & Talk* and *Choice & Backlash* on women, reported at the bottom of the table, is 0.692, which indicates that the possibility of backlash does not increase mistakes for women.<sup>33</sup> Our interpretation is that the gender difference in likelihood to make mistakes in the presence of backlash is caused by the negative effect of the treatment on men, rather than on a positive effect of the treatment on women. In fact, the uninteracted coefficient of *Choice & Backlash* indicates the the possibility of backlash impacts men's likelihood of making mistakes negatively, although the coefficient is not statistically significant ( $p\text{-value}=0.116$ ). Columns 4 to 6 of Table 4 also show that self-selected female managers are less likely to make ranking mistakes as compared to not self-selected female managers.

As a final assessment of managerial performance, we look at lost profits caused by mistakes in rank allocations. Recall that the manager receives earnings from the number of puzzles that are successfully completed by the rank A worker only. If a manager assigns rank A to a worker who ends up performing worse than the other worker, the manager suffers a loss in profit equal to the difference between the profit that he or she would have made by allocating rank A to the best performing worker (i.e., the counterfactual) and the profits that he or she actually

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<sup>32</sup>Estimates obtained for the male and female samples separately are displayed in Table A3 in the appendix.

<sup>33</sup>The p-value obtained for the linear combination "CB + Female x CB" being equal to 0, which indicates the impact of *Choice & Backlash* on women as compared to *Choice*, is equal to 0.225 (column 6).

made. Lost profits are equal to zero if the manager was not mistaken in the rank allocation, and equal to the difference in the number of puzzles solved by the two workers, multiplied by the piece rate (2 ECU), if he or she was. Results from regression analysis, displayed in columns 7 and 9 of Table 4, show no evidence of treatment or gender effects on lost profits: neither having to talk to employees when assigning ranks, or facing the possibility of receiving angry messages from low-ranked workers leads male and female managers to make decisions that ultimately result in financial losses. Note that lost profits are low on average, i.e., below 2 ECU. This is both because mistakes only happen 22 percent of the times on average, and because, when they do happen, the difference in the performances of the two workers tends to be small, i.e., less than 4 puzzles on average.

In Table 5, we compare the total profits that male and female managers made over the 5 stages of the experiment from the productivity of the workers assigned rank A. We find no evidence of significant differences by treatment and by manager gender. Table 5 also displays the difference in the total earnings of the two workers. While there are no significant differences in inequality between workers' earnings by manager gender in any of the treatments, the inequality tends to be reduced when the manager is a woman and there is the possibility of backlash.

Finally, we assess gender and treatment differences in manager leadership styles, as measured by the free-form messages that managers sent to the workers at the rank assignment stage in the *Choice & Talk* and the *Choice & Backlash* treatments. We conducted text analysis<sup>34</sup> of the 360 messages sent by managers over the 5 stages of the experiment in the *Choice & Talk* (185) and the *Choice & Backlash* (175) treatments. We identified seven primary types of messages: 1) praising the worker for past performance, e.g., “amazing work!”; 2) using motivating words, e.g., “You got this!”; 3) providing an explanation for the rank allocation, e.g., “the other worker did better”; 4) inducing competition among workers, e.g., “the other worker is catching up so keep it up”; 5) mentioning fairness, e.g., “trying to be fair”; 6) using team building words, e.g., “let’s go team!”; and 7) using cordial words like “thank you” and “sorry”.<sup>35</sup>

Table 6 shows the percentages of messages within each category sent by male and female managers to rank B workers over the five stages of the experiment, in the *Choice & Talk* and the *Choice & Backlash* treatments. In the *Choice & Talk* treatment, where workers could not send messages back, we see that men tend to explain their decisions more, and are more likely to use cordial language. Under the threat of backlash, both male and female managers explain their ranking decisions significantly more, but the gender difference remains in favour

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<sup>34</sup>The analysis was conducted by a graduate student, who had access to the messages but not to the gender of the manager or the treatment. The analysis was done in Stata, by searching for keywords in the messages, where different keywords were associated with the different message categories.

<sup>35</sup>Table A4 in the appendix provides additional examples of the messages of each kind sent by managers to workers.

of men. Moreover, a gender difference emerges in the likelihood of praising the rank B workers (significantly lower for men than for women), the likelihood of sending messages that induce competition among workers (significantly higher for men than for women). Moreover the gender difference in the likelihood of using cordial language reverses, i.e., women become more likely than men to use words such as “thank you” and “please” (p-value=0.04).<sup>36</sup> While these gender differences are in line with previous studies of differences in the language used by men and women, both in psychology (e.g., Kern et al., 2016) and economics (Timko, 2017), our treatments reveal that such gender differences are generated by differential responses to the possibility of worker backlash, rather than by innate gender differences in communication styles.<sup>37</sup>

### 3.3 Backlash and manager reaction to backlash

Over the 5 stages of the experiment, male and female managers received backlash from the rank B worker - in the form of at least one angry emoji - around 45 percent of the times. Figure 3 shows that female managers tend to receive more angry emojis. Figure A3, in the appendix, shows the percentages of male and female managers who received at least one angry emoji in each stage of the experiment. While female managers started off by being slightly more likely to receive at least one angry emoji, the difference is not statistically significant. The figures suggests that the backlash received by male and female managers differ at the intensive but not the extensive margin.

In Table 7, we estimate equation 3 of Section 2.4. In particular, we conduct regression analysis of the decision of rank B workers to send angry emojis to their manager over the 5 stages of the experiment. The dependent variable is the number of angry emojis sent, including the zeros.<sup>38</sup> We progressively add controls, and, in columns 3 and 4, we also include dummy variables capturing the type of free-form message (see Table 6) sent by the manager,<sup>39</sup> since the rank B worker sees such message immediately before he or she is given a chance to send angry emojis. The estimates confirm that female managers receive more angry emojis from

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<sup>36</sup>We do not find any gender difference, or treatment effects, in the propensity to use team building words, or to mention fairness.

<sup>37</sup>Statistics for the messages sent to rank A workers are shown in Table A5 in the appendix. The table shows evidence of no statistical significant differences in the types of messages sent by male and female managers in the *Choice & Talk*. Similarly to what we observe for the messages to the rank B workers, when worker backlash is possible, we see that men are significantly less likely than women to praise the workers, significantly more likely to explain their ranking decisions and significantly more likely to induce competition among workers. Moreover, women become more likely than men to use cordial words.

<sup>38</sup>Table A7 in the appendix replicates the analysis for the extensive margin of backlash, i.e., for workers’ propensity to send at least one angry emoji to managers. The analysis confirms that there is no gender difference in received backlash at the extensive margin.

<sup>39</sup>The dummy variables correspond to the seven message categories described in Section 3.2 and displayed in Table 6. Note that the dummies are not mutually exclusive, since the same message can for example be both praising and motivating.

rank B workers. Controlling for the type of message sent by the manager strengthens this result. Rank B workers tend to send 1 more angry emoji to a female manager than to a male manager.<sup>40</sup>

We also conduct text analysis of the 175 free-form messages that rank B workers sent to managers in the *Choice & Backlash* treatment. The most common messages can be categorized as follows: 1) thanking the manager or joking; 2) expressing approval of the ranking decision; 3) expressing commitment to work hard; 4) apologetic; 5) containing negative feelings, i.e., upset; 6) questioning the ranking decision.<sup>41</sup> We do not see significant differences in the messages sent by rank B workers, with one exception. Female managers are more likely to receive messages in which rank B workers question the ranking decision. This happens 17 percent of the times for female managers, versus 6 percent for male managers ( $p=0.032$ ).<sup>42</sup>

Overall, the analysis of worker backlash suggests that rank B workers are both more likely to verbally question the ranking decision when it comes from a female manager, and more likely to express disapproval of the choice made by a female manager by sending more angry emojis. This is despite the fact that female managers switch ranks more often and that the propensity to send angry emojis declines when workers' ranks are switched, as suggested by the positive and significant sign of the dummy indicating that the worker was not switched, i.e., he or she was rank B also in the previous round.

Finally, we examine how male and female managers react to the receipt of angry emojis in the *Choice & Backlash* treatment. We report our estimates of equation 4, presented in Section 2.4, in Table 8. The dependent variable is the decision to switch ranks and promote the rank B worker at the beginning of stage  $t$ . Since the analysis is restricted to the *Choice & Backlash* treatment, we now include the number of angry emojis that the rank B worker sent to the manager when he or she was notified of the assigned rank in the previous stage. Importantly, we interact the number of angry emojis received with the gender of the manager. While the receipt of angry emojis does not affect male managers' ranking decisions, it actually lowers the likelihood that a female manager promotes the rank B worker by 7 percentage points, everything else being equal ( $p = 0.003$ , test of the linear combination of angry emojis and its interaction with the female dummy). This indicates that the observed higher propensity of female managers to promote the rank B worker in the *Choice & Backlash* treatment as

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<sup>40</sup>An interesting question, which is however beyond the scope of our original experimental design, is whether men or women are more likely to send angry emojis to managers. We are constrained by our sample size, which comprises 70 workers in the *Choice & Backlash* treatment, of which only 30 percent are women. Keeping this in mind, the average numbers of emojis sent by men and women (including the zeros), respectively, are 1.21 and 2.31 ( $p\text{-value}=p=0.002$ ).

<sup>41</sup>Table A6 in the appendix provides examples of the messages belonging to each category.

<sup>42</sup>Table A8 in the appendix shows the percentage of messages belonging to each category. There are no significant differences in the messages sent by rank A workers to male versus female managers, which belong to three main categories: 1) thanks or jokes (56 percent), 2) approval of rank assignment (7 percent), 3) commitment to work hard (23 percent).

compared to the other *Choice* treatments is not a response to *received* angry emojis and is not due to a desire to appease employees who have expressed disapproval of the ranking decision.

We replicate the analysis for manager responsiveness to “any backlash” rather than severity of backlash in Table A9 in the appendix. The estimates shows that men’s promotion decisions do not respond to receiving one or more angry emojis. In contrast, receiving any backlash significantly reduces women’s likelihood of promoting the rank B worker by about 18 percentage points (p-value=0.068, test of the linear combination of “at least one angry emoji” and its interaction with the female dummy variable).

### 3.4 Robustness: The gender composition of the group

When recruiting subjects to participate in the experiment, we aimed to have groups of two men and one woman to simulate male-dominated environments. About 63 percent of our groups ended up being of the desired gender composition, with the remaining groups consisting primarily of two women and one man. When examining self-selection into leadership in Section 3.1, we controlled for the number of women in the group, and we clustered the standard errors at the group level. In the regression analysis, the number of women in the group did not seem to impact the decision to become the manager of the group.

In Figure A4 in Appendix, we reproduce Figure 2 for the restricted sample of groups composed of two men and one woman only. The pattern we saw for the full sample remains, with women less likely than men to self-select into leadership in the *Choice & Backlash* treatment only (p=0.009). Among women, we see a lower percentage of women wanting to be manager in the *Choice & Backlash* treatment than in the *Choice & Talk* treatment, although the difference is not statistically significant in the restricted smaller sample (p=0.123). We do not see treatment-driven significant differences among male participants.

The gender composition of the group, combined with the gender of the manager, has immediate consequences for the genders of the workers a manager is matched with. In our setting, averaging across all treatments, 77 percent of male managers are matched with one male and one female worker, versus 47 percent of female managers (p=0.000). We conduct robustness checks of managers’ decision-making, restricting the analysis to groups where both male and female managers are matched with one male and one female worker. The results obtained for rank-switching, ex-post mistakes and lost profits are displayed in Table A10 in the appendix. Even though the sample is smaller by about one third, the signs of the estimates are consistent with those obtained in Table 4, with male and female managers more willing to switch ranks when having to talk with employees, and overall no significant differences in the likelihood of ex-post mistakes and lost profits.<sup>43</sup>

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<sup>43</sup>The higher likelihood of female managers to switch ranks in the *Choice & Backlash* treatment disappears



## 4 Mechanisms: The online experiment

The primary findings generated by the laboratory experiment could be driven by women’s expectations of more backlash or by women’s higher sensitivity to any backlash, or both. Our lab experimental data do not allow us to test for the mechanisms driving the results, as we did not collect data on beliefs about backlash. This was to prevent participants from being influenced to consider the severity of potential backlash during the self-selection into leadership phase in the *Choice & Backlash* treatment, beyond what they would have naturally done. We also worried that eliciting incentivised beliefs on backlash in only one of the treatments - the one where backlash was possible - would introduce methodological differences between the treatments that would compromise the internal validity of the experiment.

In order to generate insights on the mechanisms that could drive our findings - aversion to backlash and expectation of backlash - we conducted an online experiment on Prolific in Fall 2021. We involved a sample of over 500 participants of similar age and education as our laboratory subjects. Specifically, we restricted recruitment to individuals aged between 18 and 30, either currently enrolled in college or with a college degree. The online survey had three main parts: 1) A brief section on demographics (to check that the recruitment requirements were met); 2) A set of questions used to construct the Brief Fear of Negative Evaluation (BFNE) psychology index (Rodebaugh et al., 2004); 3) The description of the laboratory experiment and the incentivized elicitation of beliefs regarding the likelihood and severity of backlash that men or women managers received.

The survey lasted 15 minute and subjects earned an average of 4.5USD, the equivalent of 18USD per hour. A total of 566 subjects participated, of which 47 percent were women. Participants were 23.5 years old on average, about half of them were enrolled in college at the time of the survey, and the vast majority (64 percent) of them held a leadership position in the past.<sup>44</sup>

### 4.1 The BFNE Index: A measure of aversion to backlash

In the online survey we include questions that allow us to construct the “straightforward” version of the Brief Fear of Negative Evaluation (BFNE) index (Rodebaugh et al., 2004; Carleton et al., 2011). This psychological index is a short version of the 30-item index originally developed by Watson and Friend (1969) and often used in personality and social psychology to measure anxiety associated with the possibility of being evaluated or judged negatively by

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in the restricted sample, but we may be underpowered to detect the treatment effect.

<sup>44</sup>Compared to the participants in our laboratory experiment, the Prolific subjects are about a year older and have experienced leadership less frequently. See Table A1 in the appendix for the descriptive statistics of our lab sample.



others.<sup>45</sup>

Figure 4 shows the distribution of the BFNE scores for men and women, with a higher value indicating higher apprehension at the prospect of negative evaluation. Women score significantly higher than men on average (29.36 versus 27.24, with  $pvalue=0.0012$ ). The two distributions are also significantly different from each other ( $pvalue=0.005$  from Kolmogorov-Smirnov test for equality of distribution functions). This indicates that, at least among individuals in this age group (18 to 30) and education level (college graduates or college enrolled), women are significantly more sensitive to the possibility of being evaluated negatively by others. This suggests that a greater aversion to backlash is a plausible driver of our laboratory findings.

## 4.2 The experiment: Predictions of backlash

A crucial objective of the online study was to elicit subjects' beliefs about the backlash that men and women managers got in the laboratory experiment. To this end, after explaining the rules of the leadership game under the *Choice & Backlash* treatment, we randomly assigned participants either to a Male Manager treatment or a Female Manager treatment. Subjects in the Male Manager treatment had to first guess the percentage of male managers that received at least one angry emoji from the rank B worker after the rank allocation (using 10-point percentage ranges, e.g. 0-10%, 11-20% etc.), and then the average number of angry emojis received by male managers (conditional on receiving at least one). Participants assigned to the Female Manager treatment answered the same questions but for female managers.<sup>46</sup> Subjects earned 0.50 USD for each correct guess.<sup>47</sup>

As a result of our between-subject design, we have both men and women being randomised to guess the backlash received by either male or female managers. Figure 5 displays our primary findings. First, women believe: 1) that female managers are more likely to receive any backlash compared to male managers (i.e., the extensive margin of backlash), and 2) that female managers receive more angry messages (i.e., the intensive margin of backlash) than men. Second, men do not hold such beliefs, i.e., they do not think that female managers receive more backlash than male managers. Third, both men and women overestimate the

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<sup>45</sup>Rodebaugh et al. (2004) and Weeks et al. (2005) have shown that the 8-straightforwardly-worded-item index is more reliable than the original more comprehensive index, which also employed reverse-scored items. See the Online Appendix for detailed information on the questions forming the BFNE index employed here.

<sup>46</sup>To simplify the setting and make belief elicitation easier, we presented the game as a one-shot game, and we incentivised beliefs based on the occurrence of backlash in stage 2 of the lab experiment, i.e., the first stage in which backlash could occur.

<sup>47</sup>We included three comprehension questions. About 7 percent of participants failed at least one comprehension question. Only 1 participant failed all three questions. In the analysis in this subsection, we exclude the subjects who failed at least one comprehension question, although the results are robust to keeping them in the sample.

occurrence of backlash by about 20 to 30 percentage points on average, and the severity of backlash by about 1 angry emoji.<sup>48</sup> Finally, while women expect women to get more backlash than men, there is no statistical significant difference in the likelihood and severity of backlash than men and women expect their own gender to receive. In fact, men expect between 70 and 80 percent of male managers, and women expect between 70 and 80 percent of female managers, to get backlash. The same applies to the severity of backlash: men and women expect male and female managers, respectively, to receive between 3 and 4 angry emojis on average.

Overall, the online experiment suggests that, while women expect women to receive more backlash than men, the impact of the backlash treatment (only) on women in the laboratory experiment - both at the self-selection into leadership and rank assignment stages - is unlikely to be driven by gender differences in expectations of backlash. Gender differences in sensitivity to backlash may be particularly important. This is in line with our findings regarding managers' reaction to receiving angry emojis in the laboratory experiment. Our analysis in Table 8 showed that men ignore the received backlash. Female managers instead react to it. Yet, they do so not by appeasing the workers who sent the emojis; rather, they become less likely to promote them (from a starting point where they do so more than men).

## 5 Conclusion

The literature on gender differences in leadership is still in its nascent stage. In this paper, we asked whether women are less likely to self-select into leadership positions that require discretionary decision-making likely resulting in backlash. By employing a controlled experiment, we were able to assess also the causal impact of the threat of backlash on leader decision-making and final outcomes.

We found strong evidence of gender differences in individuals' reaction to the possibility of backlash. While men are unaffected, women are less likely to want to assume a leadership role under the threat of backlash. No gender differences are observed when backlash is not allowed, with over 90 percent of both men and women wanting to assume the leadership role. We also find that the possibility of worker backlash leads women managers to switch worker ranks, by promoting the low-rank worker, more often, but this does not lead to losses in profits or to gender differences in final outcomes. Further analysis shows that the higher likelihood of rank-switching among women under the threat of backlash is not a reaction to received backlash, i.e., a desire to appease a worker who expressed discontent.

We also found gender differences in the messages sent by managers to workers. Female

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<sup>48</sup>See Figure A3 in the appendix for the actual extensive margin of backlash by stage, and panel b of Figure 3 for the actual intensive margin of backlash by stage.

managers are more likely to praise employees and use cordial words, while male managers are more direct and more likely to foster competition among employees. However, these differences emerge only under the possibility of worker backlash. Finally, our laboratory findings confirm previous studies on differential attitudes toward male and female leaders. In our setting, female managers receive more backlash, and are more often questioned by workers regarding their ranking decisions.

In order to assess whether gender differences in expectations of backlash and/or in aversion to any backlash are the main drivers of our laboratory findings, we conducted an online experiment on Prolific. We recruited a sample of survey respondents of comparable age and education level as our laboratory subjects. We elicited survey questions aimed at constructing the Brief Fear of Negative Evaluation psychological index. We then presented subjects with the rules of the leadership game in the *Choice & Backlash* treatment, and randomly assigned them to predict the backlash received by either male or female managers in the laboratory experiment. We find evidence of significant gender differences in the Brief Fear of Negative Evaluation index, with women being more averse to the possibility of negative judgement. Results from the experimental elicitation of beliefs about backlash show that, while women expect women to receive more backlash than men, there is no gender differences in the backlash that men and women expect their gender to receive. This suggests that expectations of backlash are unlikely to be the driving mechanism behind our primary lab findings.

In evaluating the relevance and implications of our results, it is important to note that the analysis is based on a laboratory setting where decisions are anonymous and worker disapproval takes the form of angry emojis sent via a computer terminal. In the field, backlash can take many forms. Less severe cases of backlash include non-verbal expressions such as angry emojis on social platforms, angry in-person glances, avoiding physical contact (e.g., avoiding walking in the same hall), exclusion from social activities (e.g., lunch), and the sharing of criticism with coworkers and others (often anonymously) on social platforms. Stronger forms of backlash include on-the-job sabotaging and filing formal complaints with upper-level management or the HR department. In extreme cases, backlash may escalate to verbal or physical abuse. Our experiment simulates a mild form of backlash that does not affect the manager's earnings or physical well-being in any way. Moreover, backlash comes from an individual who does not have a long work or personal history with the manager. Finally, it comes from one worker only, as opposed to an environment where manager decision-making may lead to backlash from a larger number of employees. Our finding of a gender disparity in how individuals respond to the possibility of this mild form of backlash implies that the presence of more severe forms of backlash would result in an even wider gender leadership gap. In other words, we are likely underestimating the role that the threat of backlash plays in contributing to the gender leadership gap in field settings.

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## Figures and Tables

Figure 1: Stages of the experiment

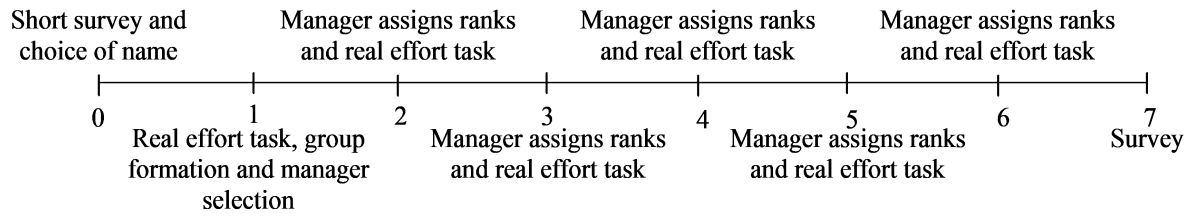
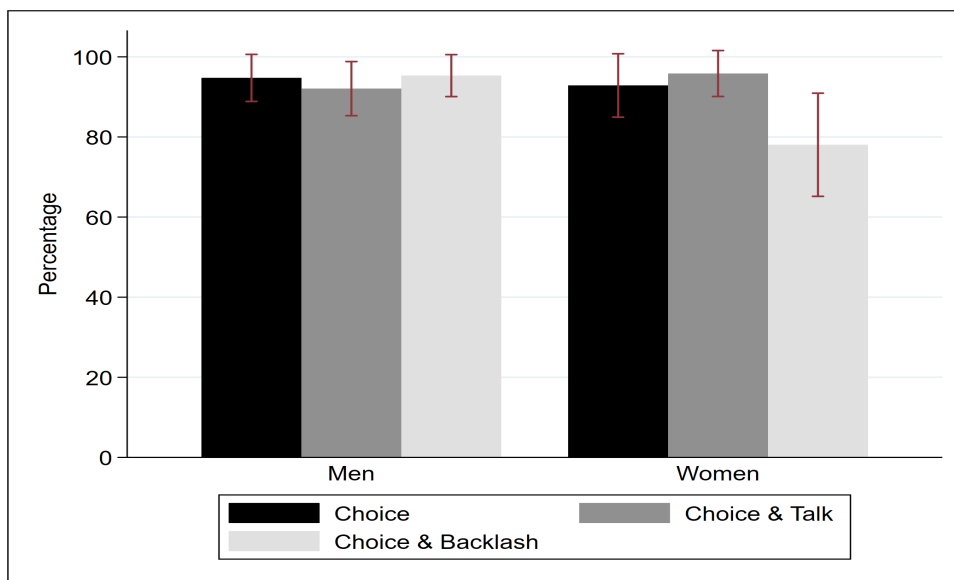
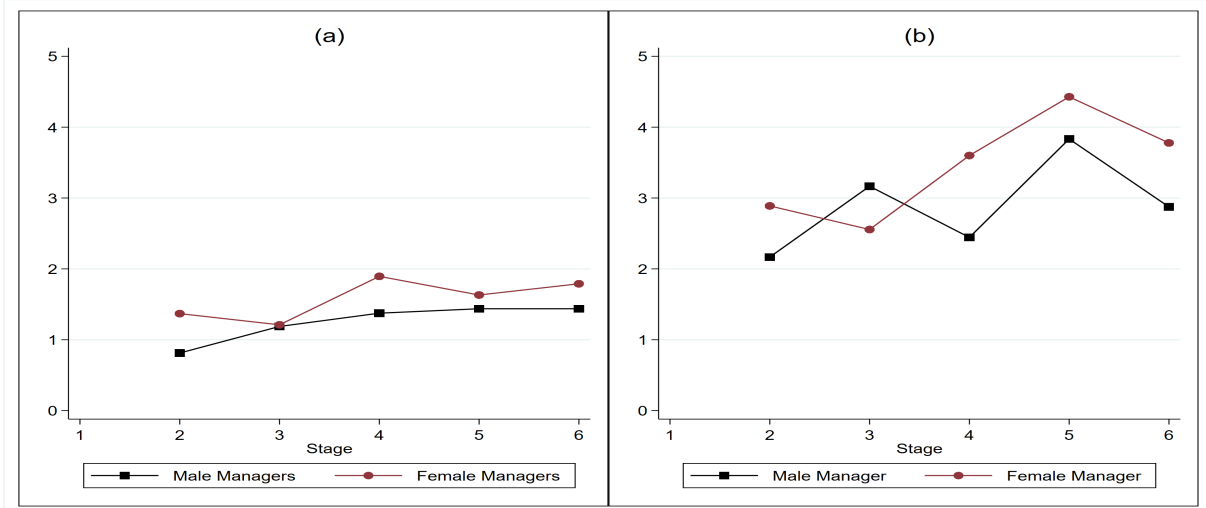


Figure 2: Self-selection into the manager position



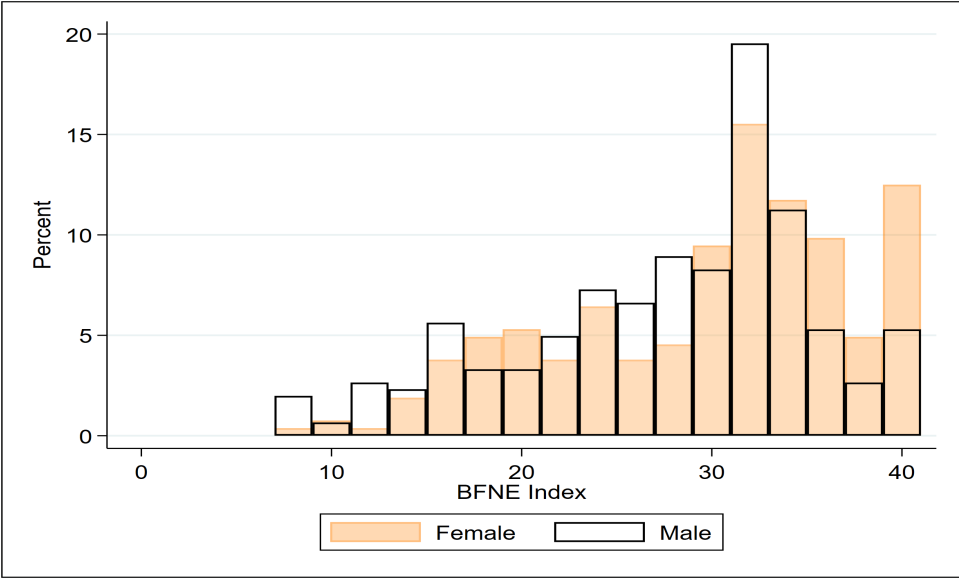
Note: The figure shows the average percentages (and 95% confidence intervals) of men and women who stated they wanted to be the manager of their group, by treatment and by gender.

Figure 3: Angry emojis sent to managers



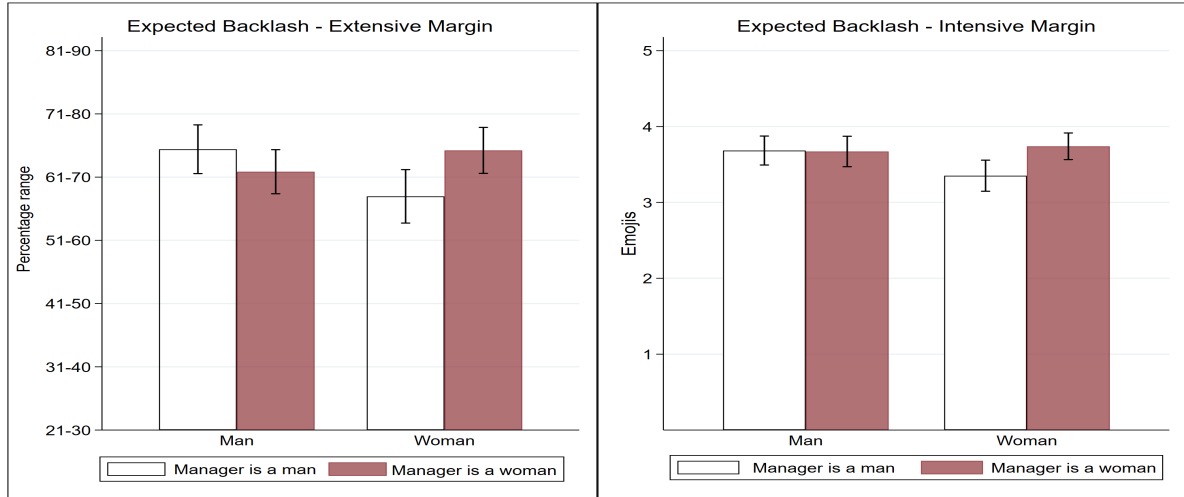
Note: Panel (a) shows the average number of angry emojis sent by rank B workers to male and female managers in the *Choice & Backlash* treatment. Panel (b) shows the emojis sent, conditional on sending at least one.

Figure 4: Online experiment: The BFNE Index



Note: The figure displays how men and women who participated in the online experiment scored in the Brief Fear of Negative Evaluation (BFNE) index. A larger value indicates higher fear of negative evaluation.

Figure 5: Online experiment: Predictions of backlash



Note: Participants in the online experiment were asked to predict the percentage of male or female managers (experimentally manipulated through a between-subject design) who received angry emojis from the rank B worker. The left panel shows the average percentage range (and 95 % confidence intervals) of man/woman managers that participants guessed received at least one angry emoji. The right panel shows the average number of angry emojis that participants (and 95 % confidence intervals) guessed man/woman managers received from rank B workers, conditional on receiving at least one.

Table 1: Groups and Treatments

	Groups	Men	Women	Total
No Choice (NC)	34	61	41	102
Choice(C)	33	57	42	99
Choice & Talk (CT)	37	63	48	111
Choice & Backlash (CB)	35	64	41	105
Total	139	245	172	417

Note: Subjects participated in one of four treatments following a between-subject design. Participants were placed in a group of three people and remained in that group for the duration of the experiment.

Table 2: Men and Women who want to be managers

	Men	Women	$H_0 : M = W$
Choice(C)	94.74	92.86	0.70
Choice & Talk (CT)	92.06	95.83	0.42
Choice & Backlash (CB)	95.31	78.05	0.01
$H_0 = C = CT$ (p-value)	0.56	0.54	.
$H_0 = C = CB$ (p-value)	0.88	0.06	.
$H_0 = CT = CB$ (p-value)	0.45	0.01	.

Note: We report the percentages of participants who wanted to be the managers of their group. P-values are generated by Chi-square tests.

Table 3: Self-selection into the manager position - Pooled sample

	Dep. Var: Wants to be manager			
	(1)	(2)	(3)	(4)
Choice&Talk (CT)	-0.002 (0.037)	-0.027 (0.055)	-0.009 (0.051)	-0.003 (0.050)
Choice&Backlash (CB)	-0.055 (0.038)	0.006 (0.047)	0.010 (0.046)	0.025 (0.046)
Female	-0.049 (0.036)	-0.019 (0.057)	-0.007 (0.055)	-0.003 (0.057)
Female x CT		0.056 (0.077)	0.045 (0.074)	0.040 (0.074)
Female x CB		-0.154 (0.092)	-0.160 (0.090)	-0.164 (0.095)
Constant	0.960 (0.032)	0.947 (0.038)	0.907 (0.131)	0.816 (0.145)
Controls	No	No	Yes	Yes <sup>+</sup>
Observations	315	315	315	315
Clusters	105	105	105	105
H <sub>0</sub> : CT = CB	0.158	0.499	0.639	0.504
H <sub>0</sub> : CT + F. x CT=0		0.548	0.455	0.464
H <sub>0</sub> : CB + F. x CB=0		0.047	0.037	0.069
H <sub>0</sub> : CT + F. x CT=CB + F. x CB		0.011	0.008	0.013

Note: Linear probability models. Robust standard errors, clustered at the group level, in parentheses. The bottom four rows report p-values generated by the corresponding hypothesis tests. Controls are: performance in the task in Stage 1, age, and STEM or economics field of study. Additional controls in column 4 are: Being native English speaker, having held a leadership position, Big5 Agreeableness index and number of women in the group. Due to a computer glitch, the additional individual characteristics in Column 4 were not collected for 12 subjects in the *Choice* treatment. We input the missing values of the Big 5 Agreeableness index with its mean, and the missing values of the past leadership indicator and the native English speaker indicator with the most common value in the non-missing sample. In Column 4, we include a dummy variable equal to 1 for the 12 subjects who have missing values for the three additional control variables.



Table 4: Manager decision-making

	Switched ranks			Ex-post mistake			Lost profit		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
No Choice (NC)	0.104 (0.067)	0.101 (0.069)	0.099 (0.070)	-0.021 (0.081)	-0.020 (0.081)	-0.037 (0.079)	0.866 (0.817)	0.852 (0.829)	0.696 (0.834)
Choice&Talk (CT)	0.172 (0.068)	0.182 (0.074)	0.183 (0.078)	0.034 (0.090)	0.015 (0.088)	-0.013 (0.088)	0.497 (0.708)	0.343 (0.700)	0.077 (0.744)
Choice&Backlash (CB)	0.149 (0.084)	0.129 (0.083)	0.126 (0.082)	-0.121 (0.076)	-0.109 (0.077)	-0.120 (0.076)	-0.318 (0.671)	-0.346 (0.657)	-0.397 (0.654)
Female	0.042 (0.069)	0.024 (0.072)	0.009 (0.074)	-0.089 (0.075)	-0.081 (0.076)	-0.089 (0.075)	-0.179 (0.695)	-0.238 (0.732)	-0.357 (0.756)
$\Delta$ Performance <sub>t-1</sub>	-0.037 (0.003)	-0.037 (0.003)	-0.037 (0.003)	-0.014 (0.003)	-0.014 (0.003)	-0.014 (0.003)	-0.129 (0.031)	-0.129 (0.031)	-0.129 (0.031)
Female x NC	0.000 (0.104)	0.008 (0.104)	0.025 (0.108)	0.181 (0.100)	0.176 (0.100)	0.188 (0.101)	0.011 (1.078)	0.020 (1.094)	0.190 (1.116)
Female x CT	-0.033 (0.101)	-0.047 (0.103)	-0.035 (0.109)	0.036 (0.109)	0.050 (0.106)	0.064 (0.107)	-0.173 (0.968)	-0.115 (0.972)	0.129 (0.992)
Female x CB	0.137 (0.117)	0.155 (0.116)	0.170 (0.117)	0.209 (0.097)	0.200 (0.099)	0.197 (0.095)	0.864 (1.018)	0.919 (1.030)	0.888 (1.037)
Constant	0.228 (0.056)	0.492 (0.130)	0.374 (0.216)	0.321 (0.074)	0.105 (0.132)	-0.047 (0.177)	1.672 (0.548)	0.784 (1.198)	-0.935 (1.666)
Observations	556	556	556	556	556	556	556	556	556
Controls	No	Yes	Yes	No	Yes	Yes	No	Yes	Yes
Stage FE	Yes	Yes	Yes+	Yes	Yes	Yes+	Yes	Yes	Yes+
Clusters	139	139	139	139	139	139	139	139	139
H <sub>0</sub> : NC = CT	0.336	0.278	0.271	0.503	0.656	0.753	0.670	0.550	0.469
H <sub>0</sub> : NC = CB	0.597	0.737	0.748	0.138	0.178	0.206	0.155	0.138	0.181
H <sub>0</sub> : CT = CB	0.799	0.536	0.511	0.047	0.097	0.157	0.262	0.318	0.517
H <sub>0</sub> : NC + Fem. × NC = 0	0.198	0.166	0.119	0.008	0.010	0.017	0.188	0.195	0.223
H <sub>0</sub> : CT + Fem. × CT = 0	0.066	0.080	0.055	0.255	0.278	0.427	0.625	0.727	0.772
H <sub>0</sub> : CB + Fem. × CB = 0	0.001	0.000	0.000	0.153	0.152	0.225	0.471	0.460	0.548
H <sub>0</sub> : CT + Fem. × CT = CB + Fem. × CB	0.077	0.087	0.095	0.781	0.691	0.692	0.738	0.606	0.667

Note: LPM in columns 1 to 6; OLS in columns 7 to 9. Robust standard errors, clustered at the manager level, in parentheses. The dependent variables are: i) a dummy equal to 1 if the Manager switched ranks (columns 1-3); ii) a dummy equal to 1 if the the worker chosen to be rank A ended up performing worse than the other worker (columns 4-6); iii) the difference between the profits that the manager would have made by assigning rank A to the best performing worker, and the profits actually made in the current Stage (columns 7-9). Lost profits are expressed in Experimental Currency Units (ECU). The analysis is restricted to Stages 3 to 6, since Stage 3 is the first stage when rank switching is allowed.  $\Delta$ performance<sub>t-1</sub> is the difference in the performances of the rank A and the rank B workers in previous Stage. Controls are: age and STEM or economics field of study. Additional controls in column 3, 6 and 9 are: native English speaker, having held a leadership position, and Big5 Agreeableness Index. Due to a computer glitch, the additional individual characteristics in Column 4 were not collected for 12 subjects in the *Choice* treatment. We input the missing values of the Big 5 Agreeableness index with its mean, and the missing values of the past leadership indicator and the native English speaker indicator with the most common value in the non-missing sample. In Column 4, we include a dummy variable equal to 1 for the 12 subjects who have missing values for the three additional control variables.

Table 5: Worker-generated Manager earnings and inequality in workers' earnings

	Total earnings generated by Rank A workers			Inequality in workers' earnings		
	Man Manager	Woman Manager	$M = W$ pvalue	Man Manager	Woman Manager	$M = W$ pvalue
No Choice (NC)	166.13	174.22	0.38	261.75	193.89	0.11
Choice(C)	161.87	177.67	0.14	246.40	278.44	0.44
Choice & Talk (CT)	177.69	179.83	0.84	176.77	212.75	0.42
Choice & Backlash (CB)	181.38	168.53	0.20	240.63	180.10	0.15
$H_0 : C = NC$	0.72	0.67	.	0.72	0.04	.
$H_0 : NC = CT$	0.31	0.53	.	0.09	0.62	.
$H_0 : NC = CB$	0.07	0.58	.	0.62	0.74	.
$H_0 : C = CT$	0.22	0.81	.	0.16	0.09	.
$H_0 : C = CB$	0.05	0.38	.	0.89	0.02	.
$H_0 : CT = CB$	0.68	0.28	.	0.20	0.40	.

Note: We report the total earnings generated by rank A workers to the manager, and the difference in the total earnings of the highest and the lowest earning worker within each group. Both total profits and differences in workers' earnings are computed over the 5 stages of the experiment. The bottom six rows report p-values generated by the corresponding hypothesis tests. All p-values are generated by tests of equality of means.

Table 6: Manager messages to rank B workers

	Praise	Motivate	Explain	Compete	Fairness	Team	Cordial
<i>Choice &amp; Talk (CT)</i>							
Male Manager	0.54	0.29	0.34	0.09	0.06	0.00	0.26
Female Manager	0.59	0.38	0.13	0.04	0.06	0.05	0.16
$H_0 : M = F(p - values)$	(0.48)	(0.22)	(0.00)	(0.16)	(0.93)	(0.07)	(0.09)
<i>Choice &amp; Backlash (CB)</i>							
Male Manager	0.29	0.23	0.54	0.23	0.07	0.03	0.20
Female Manager	0.42	0.32	0.33	0.06	0.08	0.04	0.34
$H_0 : M = F(p - values)$	(0.07)	(0.18)	(0.00)	(0.00)	(0.82)	(0.54)	(0.04)
For M							
$H_0 : CT = CB(p - values)$	(0.00)	(0.04)	(0.00)	(0.00)	(0.64)	(0.75)	(0.92)
For F							
$H_0 : CT = CB(p - values)$	(0.02)	(0.55)	(0.03)	(0.90)	(0.44)	(0.68)	(0.01)

Note: We report the percentage of messages of each type sent by male and female managers to the the rank B workers over 5 stages of the experiment in the *Choice & Talk* and the *Choice & Backlash* treatments. The analysis is based on a total of 360 messages, of which 185 in the *Choice & Talk* treatment and 175 in the *Choice & Backlash* treatment. Examples of messages in each category are provided in the online appendix. Since the categories are not mutually exclusive, the percentages do not sum up to 1. P-values are generated by Chi-square tests.

Table 7: Received Backlash

	Dep. Var: Angry emojis sent to Manager			
To Female Manager	0.823 (0.452)	0.744 (0.440)	0.920 (0.500)	0.991 (0.508)
Performance <sub>t-1</sub>	0.097 (0.040)	0.101 (0.038)	0.090 (0.039)	0.073 (0.039)
RankB <sub>t-1</sub>	0.918 (0.430)	0.777 (0.391)	1.046 (0.401)	1.051 (0.406)
Praising Message from Manager			-0.386 (0.388)	-0.419 (0.401)
Motivating Message from Manager			-0.100 (0.295)	-0.104 (0.270)
Explaining Message from Manager			-0.020 (0.360)	-0.007 (0.375)
Competing Message from Manager			-0.411 (0.459)	-0.437 (0.463)
Fairness Message from Manager			1.384 (0.625)	1.387 (0.652)
Team Building Message from Manager			0.135 (0.148)	0.155 (0.140)
Cordial Message from Manager			-0.652 (0.407)	-0.642 (0.416)
Constant	-0.982 (0.816)	-0.346 (1.616)	-0.376 (1.507)	-0.771 (2.087)
Observations	140	140	140	140
Controls	No	Yes	Yes <sup>+</sup>	Yes <sup>+</sup>
Stage FE	Yes	Yes	Yes	Yes
Manager's message type	No	No	No	Yes
Clusters	35	35	35	35

Note: OLS regressions. Robust standard errors, clustered at the group level, in parentheses. The dependent variable is the number of angry emojis, including zeros, that the rank B worker sent to his or her Manager in the *Choice & backlash* treatment. Performance<sub>t-1</sub> is the number of correct puzzles solved by the rank B worker in the previous Stage. RankB<sub>t-1</sub> is a dummy equal to 1 if the rank B worker was rank B also in the previous Stage. Controls are: age, gender and STEM or economics field of study. Additional controls in column 4 are: being a native English speaker, having held a leadership position, and Big5 Agreeableness index. In column 3 and 4, we add measures of for the type of message that the manager sent to the B worker when assigning ranks. For example of the seven not-mutually exclusive message types are presented in the online appendix.

Table 8: Manager reaction to received backlash

	Dep. Var: Switched ranks & promoted the B worker			
Female	0.182 (0.093)	0.263 (0.105)	0.234 (0.099)	0.225 (0.098)
Angry Emojis <sub>t-1</sub>	-0.055 (0.020)	-0.019 (0.024)	-0.037 (0.025)	-0.028 (0.029)
$\Delta$ Performance <sub>t-1</sub>	-0.047 (0.006)	-0.047 (0.006)	-0.045 (0.006)	-0.045 (0.006)
Female $\times$ Angry Emojis <sub>t-1</sub>		-0.061 (0.035)	-0.034 (0.033)	-0.039 (0.040)
Constant	0.622 (0.110)	0.581 (0.110)	1.284 (0.366)	1.390 (0.539)
Observations	140	140	140	140
Controls	Yes	Yes	Yes	Yes <sup>+</sup>
Stage FE	Yes	Yes	Yes	Yes
Clusters	35	35	35	35
$H_0 : \text{Emojis}_{t-1} + \text{Fem.} \times \text{Emojis}_{t-1} = 0$		0.003	0.001	0.003

Note: Linear probability models. Robust standard errors, clustered at the group level in parentheses. The dependent variable is a dummy equal to 1 if the Manager switched ranks and promoted the rank B worker at the beginning of the current Stage in the *Choice & Backlash* treatment. Angry Emojis<sub>t-1</sub> is the number of angry emojis sent by the rank B worker in the previous Stage.  $\Delta$ performance<sub>t-1</sub> is the difference in the performances of the rank A and rank B workers in the previous Stage. Controls are: age and STEM or economics field of study. Additional controls in column 4 are: being a native English speaker, having held a leadership position, and Big5 Agreeableness index.